# Input Output HMMs for modeling network dynamics

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The style of IOHMM is adapted from Prof. Craven's lectures on HMMs

# **Goals for today**

- What are Hidden Markov Models (HMMs)?
  - How do they relate to Dynamic Bayesian Networks?
- What are Input/Output HMMs (IOHMMs)?
- EM algorithm for learning IOHMMs
- Application of IOHMMs to examine regulatory network dynamics

### Motivation

- Suppose we are given time series expression profiles
- We wish to find key regulators that are associated with changes in expression levels over time
- Can we more explicitly take time into account?

#### **DREM: Dynamic Regulatory Events Miner**



Ernst et al., 2007, Mol Sys Biol

## **Recall Markov chain**

- A Markov chain is a probabilistic model for sequential observations where there is a dependency between the current and the previous state
- It is defined by a graph of possible states and a transition probability matrix defining transitions between each pair of state
- The states correspond to the possible assignments a variable can state
- One can think of a Markov chain as doing a random walk on a graph with nodes corresponding to each state

#### A three state Markov chain



# Hidden Markov Models

- Hidden Markov models are also probabilistic models used to model sequential data about a dynamical system
- At each time point the system is a hidden state that is dependent upon the previous states (history)
- The observation sequence is the output of a hidden state
- HMMs are defined by observation models and transition models

Murphy 2000

## Notation

• States are numbered from 1 to K

observed character at position t

- $x = \{x_1, \cdots, x_T\}$  Observed sequence
- $\pi = \{\pi_1, \cdots, \pi_T\}$  Hidden state sequence or path
- Transition probabilities

$$a_{kl} = P(\pi_{t+1} = l | \pi_t = k)$$

Emission probabilities: Probability of emitting symbol
 b from state k

$$e_k(b) = P(x_t = b | \pi_t = k)$$

#### What does an HMM do?

- Enables us to model observed sequences of characters generated by a hidden dynamic system
- The system can exist in a fixed number of "hidden" states
- The system *probabilistically transitions* between states and at each state it *emits* a symbol/character

# **Defining an HMM**

- States
- Emission alphabet
- Parameters
  - State transition probabilities for probabilistic transitions from state at time t to state at time t+1
  - Emission probabilities for probabilistically emitting symbols from a state

#### An HMM for an occasionally dishonest casino



What is hidden? Which dice is rolled

What is observed? Number (1-6) on the die

# Formally defining a HMM

- States
- Emission alphabet
- Parameters
  - State transition probabilities for probabilistic transitions from state at time t to state at time t+1
  - Emission probabilities for probabilistically emitting symbols from a state

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#### Recall a DBN for *p* variables and *T* time points



#### **HMM represented as a DBN**



- A DBN could be used to represent the transition probabilities more compactly.
- For example, consider the state variable to be *D*-dimensional each with *K* possible values.
  - For example we are tracking *D* objects and each object can have *K* possible settings
  - The state variable can have *K*<sup>D</sup> possible values
- An HMM will attempt to model the transition probabilities between all state combinations.
- In other words, the DBN will look fully connected.

DBN

Kevin Murphy, 2000

#### DBN version of the occasional dishonest casino



## **Three important questions in HMMs**

• What is the probability of a sequence from an HMM?

- Forward algorithm

• What is the most likely sequence of states for generating a sequence of observations

– Viterbi algorithm

- How can we learn an HMM from a set of sequences?
  - Forward-backward or Baum-Welch (an EM algorithm)

#### Computing the probability of a sequence from an HMM

$$P(x_{1}, \cdots, x_{T}, \pi_{1} \cdots, \pi_{T}) = \prod_{t=1}^{T} e_{\pi_{t}}(x_{t})a_{\pi_{t}\pi_{t+1}}$$
Initial transition
$$\lim_{t \to \infty} \operatorname{Emitting symbol} x_{t}$$
State transition between consecutive time points

# Computing the probability of a sequence from an HMM

- But we don't know what the sequence of states (path) is
- So we need to sum over all paths
- The probability over *all* paths is:

$$P(x_1, \cdots, x_T) = \sum_{\substack{\pi_1 \cdots, \pi_T \\ \text{Sum over all paths}}} a_{0\pi_1} \prod_{t=1}^T e_{\pi_t}(x_t) a_{\pi_t \pi_{t+1}}$$

- The forward algorithm gives an efficient way to compute this probability
- It is based on the concept of dynamic programming

# **Goals for today**

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#### Input output Hidden Markov Models (IOHMM)

- As in the HMM we have
  - States, emissions and transitions
- In addition we have a sequence of inputs
  - The transitions and emissions can depend on inputs  $(u_1, ..., u_T)$
- In a way, IOHMMs map inputs to outputs

   This is different from HMMs
- HMMs aim to define  $P(x_1..x_T)$  while IOHMMs define  $P(x_1..x_T|u_1..u_T)$

Bengio & Frasconi, IEEE Trans on Neural Networks 1996

#### Input output Hidden Markov Models (IOHMM)



# Formally defining an IOHMM

- The set of *K* hidden states
- Emission characters/symbols/values
- Transition probabilities conditioned on the input

- Unlike HMMS where we had  $a_{kl} = P(\pi_{t+1} = l | \pi_t = k)$ - Here we have  $a_{kl} = P(\pi_{t+1} = l | \pi_t = k, u_{t+1})$ 

• Similarly for emission probabilities on the input and state

$$e_k(x_t) = P(x_t | \pi_t = k, u_t)$$

## Three important questions in IOHMMs

- What is the probability of a sequence from an IOHMM?
  - Forward algorithm
- What is the most likely sequence of states for generating a sequence of observations

– Viterbi algorithm

How can we learn an IOHMM from a set of sequences?

- Forward-backward algorithm (an EM algorithm)

#### Computing the probability of a sequence from an IOHMM

 $P(x_1,\cdots,x_T,\pi_1,\cdots,\pi_T|u_1,\cdots,u_T)$ 



#### As in the case of HMMs

• We would need to sum over the possible state configurations

$$P(x_1, \cdots, x_T | u_1 \cdots, u_T) = \sum_{\pi_1, \cdots, \pi_T} \prod_{t=1}^T P(x_t | \pi_t, u_t) P(\pi_t | \pi_{t-1}, u_t)$$

Sum over all paths

• We will use the forward algorithm for this problem

# How likely is a given sequence: Forward algorithm

• Define  $f_k(t)$  as the probability of observing  $x_1, \cdots, x_t$  and ending in state k at time t given inputs  $u_1..u_t$ 

$$f_k(t) = P(x_1, \cdots, x_t, \pi_t = k | u_1, \cdots, u_t)$$

• This can be written as follows

$$f_k(t+1) = P(x_{t+1}|\pi_{t+1}, u_{t+1}) \sum_{\substack{l=1\\k}}^{K} f_l(t) P(\pi_{t+1} = k|\pi_t = l, u_{t+1})$$
$$f_k(t+1) = e_k(x_{t+1}) \sum_{\substack{l=1\\l=1}}^{K} f_l(t) a_{lk}$$

#### **Steps of the Forward algorithm**

Initialization

$$f_k(1) = e_k(x_1)P(\pi_1 = k|u_1)$$

• Recursion: for t=2 to T

$$f_k(t) = e_k(x_t) \sum_l a_{lk} f_l(t-1)$$
• Termination

 $P(x_1, \cdots, x_T | u_1, \cdots, u_T) = \sum_{l=1}^K f_l(T)$ 

## Working through an example

- Suppose we are able to measure three reporter molecules whose values are dependent upon input chemical stimulus and whether one of four possible hidden pathways are triggered.
- Chemical stimulus: {0,1}
- Hidden Pathways: {A, B, C, D}
- Reporter molecules: {r<sub>1</sub>, r<sub>2</sub>, r<sub>3</sub>}
- Given a sequence of reporter molecule measurements, and chemical stimuli, infer which hidden pathway was likely triggered

#### Mapping to an IOHMM



$$\pi_t \in \{A, B, C, D\}$$
$$u_t \in \{0, 1\}$$
$$x_t \in \{r_1, r_2, r_3\}$$

We need to specify three CPTs

$$P(\pi_1|u_1)$$
$$P(\pi_t|\pi_{t-1}, u_t)$$
$$P(x_t|\pi_t, u_t)$$

#### The CPTs that we will use



	$x_t$							
			r <sub>1</sub>	r <sub>2</sub>	r <sub>3</sub>			
	0	Α	0.8	0.1	0.1			
	0	В	0.2	0.6	0.2			
$\pi_t, u_t$	0	С	0.25	0.5	0.25			
	0	D	0.2	0.2	0.6			
	1	Α	0.2	0.6	0.2			
	1	В	0.25	0.5	0.25			
	1	С	0.2	0.2	0.6			
	1	D	0.5	0.25	0.25			

			А	В	С	D
	0	Α	0.6	0.2	0.2	0
	0	В	0.2	0.6	0	0.2
+	0	С	0.1	0	0.8	0.1
$\pi_t, u_t$	0	D	0	0.25	0.25	0.5
	1	Α	0.8	0.1	0.1	0
	1	В	0.8	0.1	0	0.1
	1	С	0.1	0	0.8	0.1
	1	D	0	0.1	0.8	0.1

 $\pi_{t+1}$ 

Suppose we observed the following sequences Input:  $0\ 1\ 1\ 0$ Output:  $r_1\ r_1\ r_2\ r_3$ 

How likely is this observation from our IOHMM?

# Transition probabilities encode some independencies

	$\pi_{t+1}$							
			А	В	С	D		
	0	Α	0.6	0.2	0.2	0		
	0	В	0.2	0.6	0	0.2		
-	0	С	0.1	0	0.8	0.1		
	0	D	0	0.25	0.25	0.5		
2	1	Α	0.8	0.1	0.1	0		
	1	В	0.8	0.1	0	0.1		
	1	С	0.1	0	0.8	0.1		
	1	D	0	0.1	0.8	0.1		



 $\pi_t, u_{t+1}$ 

## Applying the forward algorithm

Inpu	t:	0	1	1	0	
Out	put:	r <sub>1</sub>	$r_1$	r <sub>2</sub>	r <sub>3</sub>	
		1	2	3	4	
	Α					
	В					
	С					
	D					
		$f_B($	$(3) = f_B$	= P $(2)I$	$(r_2 _{P})$	$B, 1) * (f_A(2)P(B A, 1) +  B, 1) + f_D(2)P(D C, 1))$

#### Result of applying the forward algorithm

Input:	0	1	1	0	
Output:	r <sub>1</sub>	r <sub>1</sub>	r <sub>2</sub>	r <sub>3</sub>	
	1	2	3	4	
Α	0.4	0.08	0.04488	0.0033861	
В	0.1	0.0125	0.033125	0.02186082 5	
С	0	0.008	0.00308	0.00287906 25	
D	0	0.01	0.0007625	0.00438855	

 $P(r_1, r_1, r_2, r_3|0, 1, 1, 0) = f_A(4) + f_B(4) + f_C(4) + f_D(4) = 0.0325$ 

# Learning an IOHMM from data

• Given *J* paired sequences

$$\{({m x}_{1:T_1}, {m u}_{1:T_1}), ..., ({m x}_{1:T_J}, {m u}_{1:T_J})\}$$

- Parameter estimation:
  - Learn the transition and emission probability distributions
  - This is very similar to what is done in HMMs
- Structure learning:
  - Learn the number of states and the dependencies among the states
  - Because states are hidden variables and we do not how many there are, this adds another level of complexity in learning
  - We will first assume that we know the number of states

# The expectation maximization algorithm

- Expectation Maximization (EM) is a widely used when there are hidden variables
- It is an iterative algorithm that maximizes the likelihood of the data
- Each iteration is made up of two steps
  - Expectation step (E): estimate the expected values of hidden variables given the data and previous parameter settings
  - Maximization step (M): estimate the parameters using the expected counts

#### Learning without hidden information

• Transition probabilities

$$\begin{aligned} a_{kl} &= P(\pi_t = l | \pi_{t-1} = k, u_t = p) \text{Number of transitions from} \\ &= \frac{n_{k \to l} | u_t = p}{\sum_{l'} n_{k \to l'} | u_t = p} \end{aligned}$$

Emission probabilities

$$\begin{split} e_k(c) &= P(x_t = c | \pi_t = k, u_t = p) \\ &= \frac{n_{k,c} | u_t = p}{\sum_{c'} n_{k,c'} | u_t = p} \end{split} \text{Number of times $c$ is emitted from $k$ given input $p$} \end{split}$$

#### The expectation step

 We need to know the probability of the symbol at t being produced by state i, given the entire observation and input sequence u<sub>1:T</sub>, x<sub>1:T</sub>

$$P(\pi_t = k | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T})$$

 We also need to know the probability of observations at t and (t+1) being produced by state i, and l respectively given sequence x

$$P(\pi_t = i, \pi_{t-1} = j | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T})$$

• Given these we can compute our expected counts for state transitions, character emissions

Bengio & Frasconi, IEEE Trans on Neural Networks 1996

# Computing $P(\pi_t = k | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T})$

 First we compute the probability of the entire observed sequence with the *t<sup>th</sup>* symbol being generated by state k

$$P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})$$

• Then our quantity of interest is computed as  $P(\pi_t = k | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T}) = \frac{P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}{P(\boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}$ 

Obtained from the forward algorithm

# Computing $P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})$

• To compute

$$P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})$$

• We need the forward and backward algorithm  $= P(\boldsymbol{x}_{1:t}, \pi_t = k | \boldsymbol{u}_{1:t}) P(\boldsymbol{x}_{t+1:T} | \pi_t = k, \boldsymbol{x}_{1:t}, \boldsymbol{u}_{1:T})$   $= P(\boldsymbol{x}_{1:t}, \pi_t = k | \boldsymbol{u}_{1:t}) P(\boldsymbol{x}_{t+1:T} | \pi_t = k, \boldsymbol{u}_{t:T})$ Forward algorithm  $f_k(t)$  Backward algorithm  $b_k(t)$ 

#### Steps of the backward algorithm

- Initialization (t=T)  $b_k(t) = 1$
- Recursion (t=T-1 to 1)  $b_k(t) = \sum_l a_{kl} e_l(x_{t+1}) b_l(t+1)$

# Trying out an example with backward algorithm

- Again assume we have the same CPTs as those associated with the forward algorithm demo
- Assume we observe the following

Input: 0 1 1Output:  $r_1 r_2 r_2$ 

• What are computations for the backward algorithm?

#### Results from applying the backward algorithm

 $\begin{array}{c} 0 \ 1 \ 1 \\ \text{Output:} \ r_1 \ r_2 \ r_2 \end{array}$ 

	1	2	3
Α		0.19	1
В		0.09	1
С			1
D			1

 $\pi_{t+1}$ 

		Α	В	С	D
0	Α	0.6	0.2	0.2	0
0	В	0.2	0.6	0	0.2
0	С	0.1	0	0.8	0.1
0	D	0	0.25	0.25	0.5
1	Α	0.8	0.1	0.1	0
1	В	0.8	0.1	0	0.1
1	С	0.1	0	0.8	0.1
1	D	0	0.1	0.8	0.1
	0 0 0 1 1 1 1 1	0       A         0       B         0       C         0       D         1       A         1       B         1       C         1       D         1       D         1       D         1       D         1       D	A         O       A         O       A         O       A         O       B       0.2         O       C       0.1         O       D       0         O       D       0         1       A       0.8         1       B       0.8         1       C       0.1         1       D       0	AB0A0.60.20B0.20.60C0.100C0.100D00.251A0.80.11B0.80.11C0.101D00.1	ABC0A0.60.20.20B0.20.600C0.100.80D00.250.251A0.80.10.11B0.80.101C0.100.81D00.10.8

 $b_{B}(2) = P(A|B, 1)P(r_{2}|A, 1)b_{A}(3) + P(B|B, 1)P(r_{2}|B, 1)b_{B}(3) + P(D|B, 1)P(r_{2}|D, 1)b_{D}(3)$ 

# Computing $P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})$

 Using the forward and backward variables, this is computed as

$$P(\pi_t = k | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T}) = \frac{P(\pi_t = k, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}{P(\boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}$$

$$P(\pi_t = k | \boldsymbol{u}_{1:T}, \boldsymbol{x}_{1:T}) = \frac{f_k(t)b_k(t)}{P(\boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}$$

# Computing $P(\pi_t = i, \pi_{t-1} = j | u_{1:T}, x_{1:T})$

 This is the probability of symbols at t and t+1 emitted from states k and l given the entire observed and sequence x<sub>1:T</sub> and input

sequence 
$$u_{1:T}$$
  
=  $\frac{P(\pi_t = i, \pi_{t-1} = j, \boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}{P(\boldsymbol{x}_{1:T} | \boldsymbol{u}_{1:T})}$ 

$$=\frac{f_{j}(t-1)a_{ji}e_{i}(x_{t})b_{i}(t)}{P(\boldsymbol{x}_{1:T}|\boldsymbol{u}_{1:T})}$$

## Putting it all together

• Assume we are given *J* training instances

$$\{(\boldsymbol{x}_{1:T_1}, \boldsymbol{u}_{1:T_1}), ..., (\boldsymbol{x}_{1:T_J}, \boldsymbol{u}_{1:T_J})\}$$

- Expectation step
  - Using current parameter values compute for each  $(m{x}_{1:T_J},m{u}_{1:T_j})$ 
    - Apply the forward and backward algorithms
    - Compute
      - expected number of transitions between all pairs of states
      - expected number of emissions for all states
- Maximization step
  - Using current expected counts
    - Compute the transition and emission probabilities

#### **Baum-Welch one iteration**

• Let's assume we have J=2 training instances

$$\{(0,1,1),(r_1,r_2,r_2)\}\$$

- Each training example will contribute to the expected counts of transition and emission probabilities
- Expectation step:
  - Compute the forward and backward variables for both training samples, for all time points

#### **Baum-Welch one iteration M step**

• Suppose we are updating the transition probability of A to B given input u=1  $\{(0,1,1),(r_1,r_2,r_2)\}$  $\{(1,0,0)(r_2,r_1,r_1)\}$ 

 $f_A(1)a_{AB}e_B(r_2)b_B(2) + f_A(2)a_{AB}e_B(r_2)b_B(3)$ 

$$P(r_1, r_2, r_2 | 0, 1, 1)$$

Contribution from sample 1

Sample 2 will not contribute as there is no relevant configuration

#### **Baum-Welch one iteration M step**

 Suppose we are updating the expected counts for observing r<sub>2</sub> from state B given input *u=1*

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#### **Bifurcation events**

 Bifurcation events occur when sets of genes that have roughly the same expression level up until some time point diverge



# **Dynamic Regulatory Events Miner (DREM)**

- Given
  - a gene expression time course
  - Static TF binding data or signaling networks
- Do
  - Identifies important regulators for interesting temporal changes
- DREM is suited for short time courses
- DREM is based on an Input-Output HMM

Ernst et al., 2007 Mol Sys Biol

#### **DREM key idea**



Ernst et al., 2007, Mol Sys Biol

## **IOHMM model in DREM**

- The output distributions were modeled as Gaussians
  - Enabled modeling continuous expression values
- State transitions depended on static input and the current state
  - A binary classifier was trained in the M step for each state with two children, to discriminate between genes assigned to the bifurcating states

#### Defining the transition probability in DREM

- DREM uses a binary classifier (logistic regression) to define transition probabilities
- Assume we are state *h*, which has two child states *a* and *b*

$$P(x_{t+1} = a | x_t = h, \mathbf{u}^i) = \frac{1}{1 + \exp(-\beta_0^h - \sum_f \beta_f^h \mathbf{u}^i(f))}$$

Input associated with the  $i^{th}$  gene: collection of binding sites on gene i's promoter

State-specific parameters

# Results

• Application of DREM to yeast expression data

- Amino acid (AA) starvation

– One time point ChIP binding in AA starvation

- Analysis of condition-specific binding
- Application to multiple stress and normal conditions

# DREM application in yeast amino acid starvation



DREM identified 11 paths, and associated important AA related TFs for each split

#### **Does condition non-specific data help?**



Yes, adding additional non-condition specific data helped explain more splits or found more TFs per split

# Validation of INO4 binding

- INO4 was a novel prediction by the method
- Using a small scale experiment, test binding in 4 gene promoters after AA starvation
- Measure genome-wide binding profile of INO4 in AA starvation and SCD and compare relative binding

# Validation of INO4 binding



INO4 occupancy is much higher in AA starvation compared to normal (SCD)



More genes are bound genome-wide in AA starvation



Stronger binding in AA starvation of genes in this path



# **Does integration help?**

 Randomize ChIP data and ask if enriched TFs with paths were identified

- Fewer TFs were identified

- Compare IOHMM vs HMM
  - Lesser enrichment of Gene Ontology processes in HMMs paths compared to IOHMMs

## Take away points

- Network dynamics can be defined in multiple ways
- Skeleton network-based approaches
  - + The universe of networks is fixed, nodes become on or off
  - + Simple to implement, and does not need lot of data
  - + No assumption of how the network changes over time
  - No model of how the network changes over time
  - Requires the skeleton network to be complete
- Dynamic Bayesian network
  - + Can learn new edges
  - + Describes how the system transitions from one state to another
  - + Can incorporate prior knowledge
  - Assumes that the dependency between t-1 and t is the same for all time points
  - Requires sufficient number of timepoints
- IOHMMS (DREM approach)
  - + Integrates static TF-DNA and dynamic gene expression responses
  - + Works at the level of groups of genes
  - + Focus on bifurcation points in the time course
  - Tree structure might be restrictive (although possible extensions are discussed)
  - Depends upon the completeness of the TF binding data